

- 1) Counting with no background.
 - a) "inverse" problem, $1/\sqrt{N}$
 - b) Bayesian View
 - c) Frequentist or Classical View.
- 2) Simple inclusion of background.
 - a) Relative error degrades
 - b) how do you estimate background?
- 3) Limits + discovery
 - a) Look for signal on top of expected background.
 - b) When seems like no excess over background, what do you do?
 - c) How do you discover, when it seems like "too many" events for background?

1) No Background. Somehow, "Gedanken"

Run your experiment, cuts, etc.

See, say, 25 events.

Feldman-Cousins calls this $n_0 = 25$
notation.

What you know (with scant error) —

you saw 25 events. Not much of an error in that. But... who cares?

What everyone wants to know: μ_+ , (FC notation), the "true" value of the mean # of events you expected. Very crudely, you should know —

Symmetric $\approx 68\%$ chance $25 - \sqrt{25} < \mu_+ < 25 + \sqrt{25}$
 $20 < \mu_+ < 30$

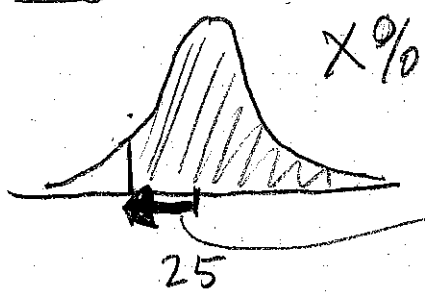


$\approx 95\%$ or $\mu_+ = 25 \pm 5$ "1 σ "

chance $25 - 2\sqrt{25} < \mu_+ < 25 + 2\sqrt{25}$

"Double Sided" $15 < \mu_+ < 35$

Single Sided :



90% only 1.28 σ !

$25 - 1.28 \cdot \sqrt{5} = 22.1$

95% — 1.64 $\sigma = 16.8$

Bayesian, Classical

$$P(n_0; \mu_+) \approx \frac{1}{\sqrt{2\pi\mu_+}} e^{-\frac{(n_0 - \mu_+)^2}{2\mu_+}}$$

deduce \nearrow \nearrow known

limiting case of $\approx \frac{\mu_+^{n_0}}{n_0!} e^{-\mu_+}$

but everyone wants to know $P(\mu_+; n_0)$

deduce \nearrow known

Bayesian: OK, get to work, but: flaws!

Classical/Frequentist: who cares what everyone wants, do what is right
Think different.

Bayes Theorem

$$P(A+B) = \underbrace{P(A; B)}_{\text{probability of A, given B}} \underbrace{P(B)}_{\text{prob of B}} = \underbrace{P(B; A)}_{\text{prob B, given A}} \underbrace{P(A)}_{\text{prob of A}}$$

Simplest case: $P(A; B) = P(A)$
statistical independence. don't care about B

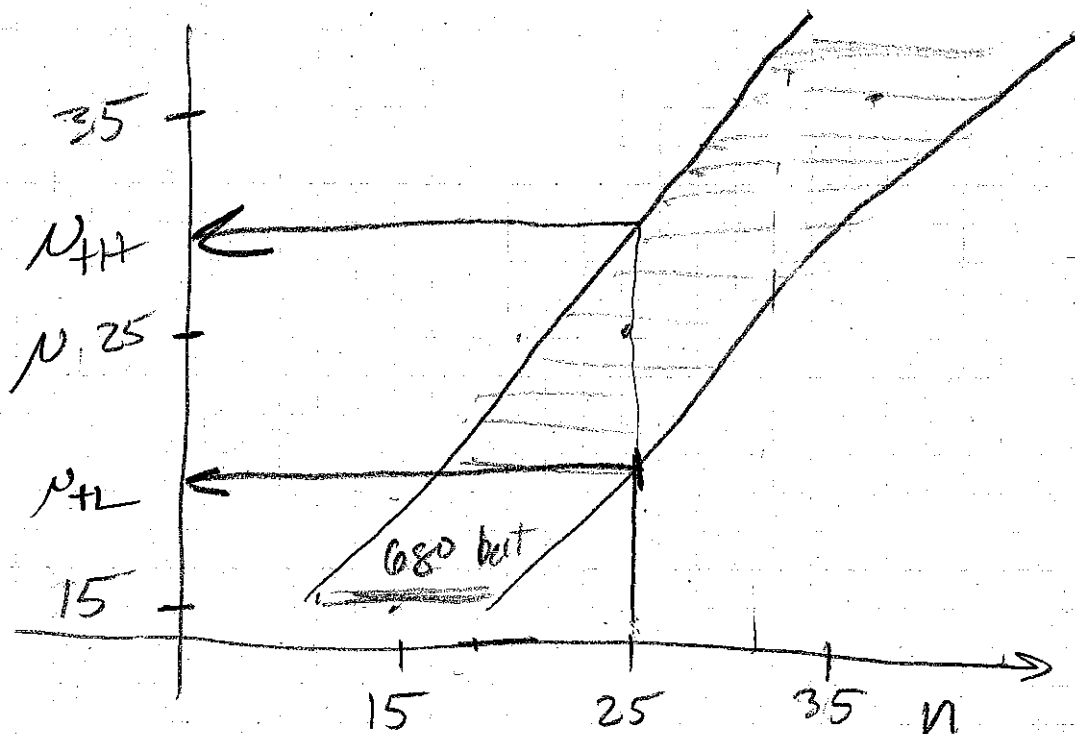
$P(B; A) = P(B)$
don't care about A.

$$P(A+B) = P(A)P(B) = P(B)P(A)$$

OK:

$$\underbrace{P(\mu_+; n_0)}_{\text{WANT}} \underbrace{P(n_0)}_{\text{"just a constant" no is known!}} = \underbrace{P(n_0; \mu_+)}_{\text{Poisson}} \underbrace{P(\mu_+)}_{\text{BELIEF "prior density" like: } \mu_+ > 0}$$

Visualization ("Neyman" Construction)



Advantage: very "experimental"

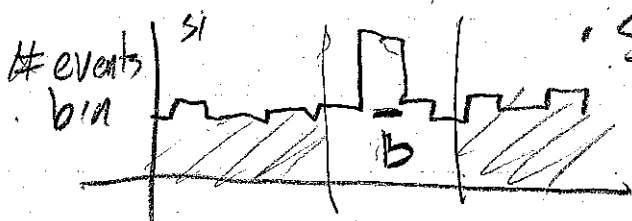
Me: Bayes when it feels good
Classical

Background: confounding

"Somehow" know it (must estimate errors)

$\mu_+ = 25$, $b = 50,000 \leftarrow$ AVERAGE

How? \uparrow
• Beam On/off
• Sidebands



M. Higgs

Roudely: Now measure $75 \pm \sqrt{75}$

$$75 \pm 8.7$$

but 50,000 of those are background
(on average)

would estimate $\mu_+ = 25 \pm 8.7$

Relative error
now...

$$\frac{8.7}{25} \approx 35\%$$

$$\rightarrow \text{not } \sqrt{75+25} = 10$$

↑
bigger than
5 because
background fluctuates

was: $\frac{5}{25} \approx 20\%$

Bad, but not as bad
maybe as you expected (2B/1 sig)
50/25

95% discovery (1)



$$50 + 1.64 \cdot \sqrt{50}$$

$$\approx 62$$

(2) 5%



$$62.77$$

which?

(don't argue,
do a better
experiment!)

→ Limitsmanship!

THE RUB

→ In previous experiment, expect 50 bkgd, want to measure an unknown signal, maybe now not 25, but small.

→ Suppose you see 35 events.

→ what do you conclude?

(A) Is $\mu = 35 - 50 = -15$?

→ "central value is!"

→ $\mu < 0$ "non physical"

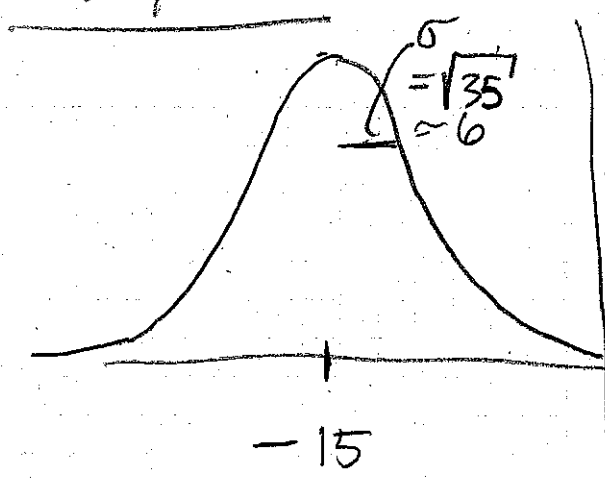
(B) Is b really equal to 50?

worry { $\frac{35-50}{\sqrt{50}} \approx \frac{-15}{7} \sim -2\sigma$

"could happen, ~ 1 in 40"

Dealing with this situation is what causes interminable discussions between Bayesians and Classical's.

Bayesian



Prob of μ_t
 $P(n_s; \mu_t, b)$

small \downarrow

n_s

Prior Belief $P(\mu_t)$

0

$\mu_t > 0$ physical

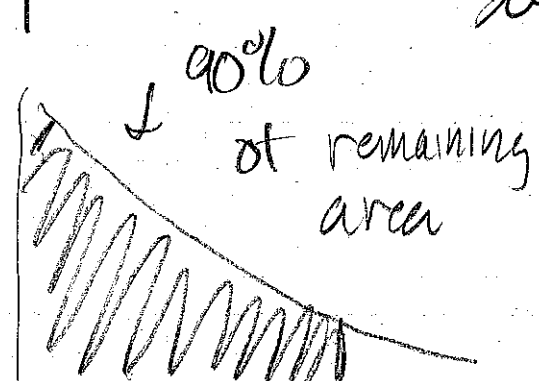
μ_t

$$P(\mu_t) P(n_s; \mu_t, b) = P(\mu_t; n_s, b)$$

UGLY!

μ_t

Blow Up \uparrow

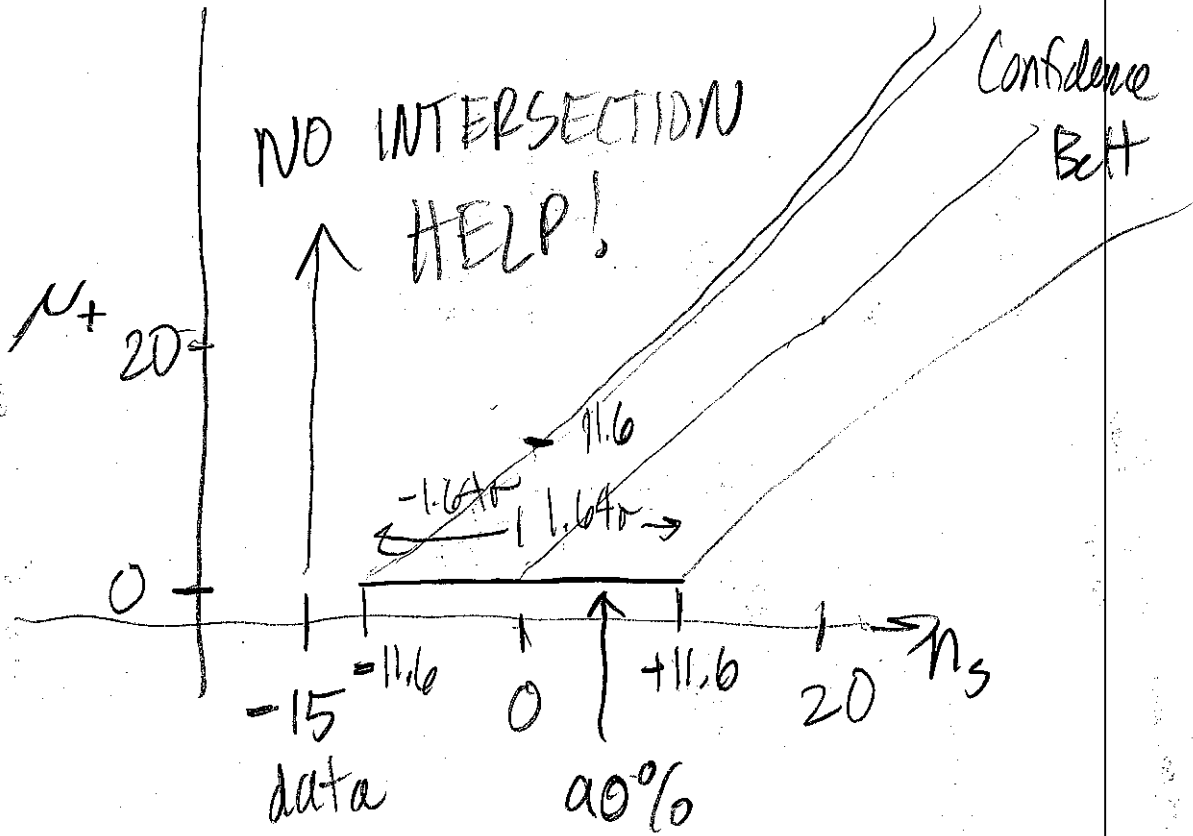


NOT EVEN IN F-C (anti-Bayesian!)

Another way?

μ_t 90% upper limit

Another way? Yes, **Frequentist**

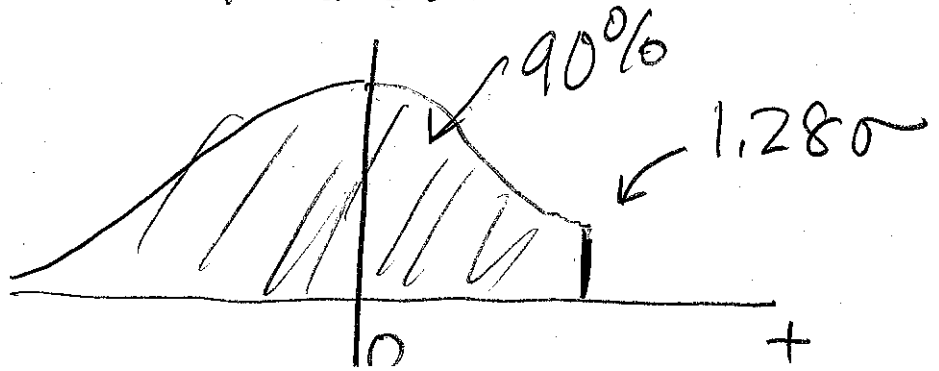


determined by...
Background!

$$1.64 \times \sqrt{50} = 11.6$$

Doesn't work either...

WAIT... UPPER LIMIT CONFIDENCE BELT ACTUALLY DESIRED...



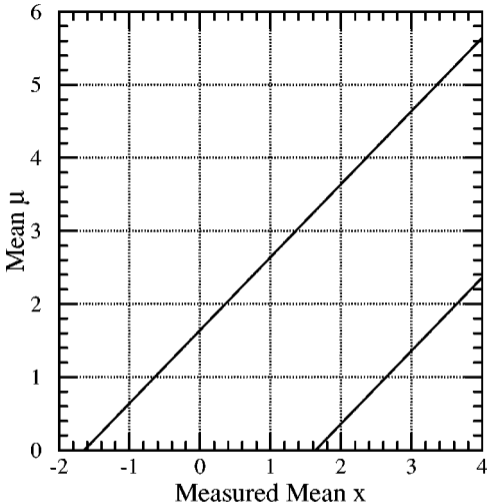


FIG. 3. Standard confidence belt for 90% C.L. central confidence intervals for the mean of a Gaussian, in units of the rms deviation.

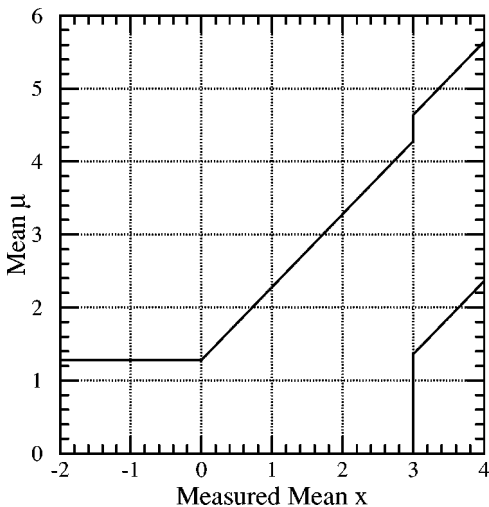


FIG. 4. Plot of confidence belts implicitly used for 90% C.L. confidence intervals (vertical intervals between the belts) quoted by flip-flopping physicist X, described in the text. They are not valid confidence belts, since they can cover the true value at a frequency less than the stated confidence level. For $1.36 < \mu < 4.28$, the coverage (probability contained in the horizontal acceptance interval) is 85%.

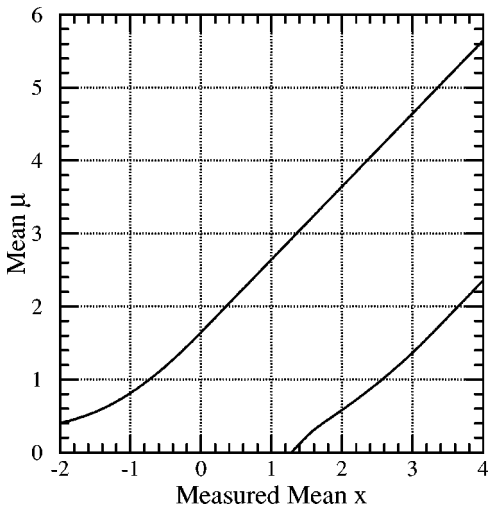


FIG. 10. Plot of our 90% confidence intervals for the mean of a Gaussian, constrained to be non-negative, described in the text.